# BAYESIAN APPROACH TO TIME DELAY ESTIMATION

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#### INTRODUCTION



Image Credit: NASA/JPL-Caltech

### INTRODUCTION



Light rays are bent by a strong gravitational field of a lensed galaxy.

- Each route has different length.
- Different arrival times of light rays

Why time delay?

- Mass structure of the lens galaxy
- ► Cosmological parameters, e.g., Hubble constant, H<sub>0</sub>

## DATA AND DIFFICULTIES



Data comprise of two time series with measurement errors.

- Observation times  $\mathbf{t}' \equiv \{t_1, t_2, \dots, t_n\}$
- Observed magnitudes  $\mathbf{x}(\mathbf{t})' \equiv \{x(t_1), x(t_2), \dots, x(t_n)\}$ , and  $\mathbf{y}(\mathbf{t})$
- Measurement errors (se)  $\delta(\mathbf{t})' \equiv \{\delta(t_1), \delta(t_2), \dots, \delta(t_n)\}$  and  $\eta(\mathbf{t})$

## DATA AND DIFFICULTIES



Some difficulties occur in estimating the time delay

- 1. Irregular observation times (: weather conditions)
- 2. Seasonal gaps (:: rotation of the earth)
- 3. Magnitude shift (:: different gravitational potentials)
- 4. Measurement errors

Our job is to estimate time delay (shift in x-axis) between two time series.

## BAYESIAN APPROACH: MOTIVATION

Grid optimization methods are dominating this field!

- eg. Cross-correlation method
  - 1. Shift one light curve by  $\Delta$  in x-axis
  - 2. Calculate  $r_{\Delta}$ , sample cross-correlation function
  - 3. Find  $\Delta$  that maximizes  $r_{\Delta}$  on the grid of  $\Delta$
- $SE(\hat{\Delta})$  by computationally expensive repeated sampling procedure
- Grid of  $\Delta$  ( $\neq$  the whole space of  $\Delta$ )

Non-grid-based Bayesian approach

- Principled way of model construction: likelihood-based
- Computational efficiency: simple and fast posterior sampling scheme
- The whole space of Δ

### BAYESIAN APPROACH: STATE-SPACE MODELING



- ► Assumption 1: ∃ unobserved underlying processes representing the true magnitudes in continuous time (red and blue dashed curves)
- ▶  $X(t)' = (X(t_1), X(t_2), ..., X(t_n))$  and Y(t), values on curves at t
- Assumption 2:  $\mathbf{Y}(\mathbf{t}) = \mathbf{X}(\mathbf{t} \Delta) + c$  (Harva, 2006)

### BAYESIAN APPROACH: LIKELIHOOD

Independent Gaussian measurement errors

- $x(t_j)|X(t_j) \sim \mathcal{N}[X(t_j), \delta^2(t_j)]$
- ►  $y(t_j)|Y(t_j) \sim \mathcal{N}[Y(t_j), \eta^2(t_j)]$  $y(t_j)|X(t_j - \Delta) + c, \Delta, c \sim \mathcal{N}[X(t_j - \Delta) + c, \eta^2(t_j)].$

Likelihood function

- Suppose  $\mathbf{t}^* = sort(t_1, t_2, \ldots, t_n, t_1 \Delta, t_2 \Delta, \ldots, t_n \Delta)$
- $\blacktriangleright L(\mathbf{X}(\mathbf{t^*}), \Delta, c) \propto \prod_{j=1}^n p(x(t_j) | \mathbf{X}(t_j)) \cdot p(y(t_j) c | \mathbf{X}(t_j \Delta), \Delta, c)$

### BAYESIAN APPROACH: PRIOR

► Ornstein-Uhlenbeck process for X(·) (Kelly et al., 2009)

- $\blacktriangleright$  Intrinsic variability of quasar  $\rightarrow$  stochastic process in continuous time
- Easy way to sample true values at irregularly-spaced times
- $dX(t) = -\frac{1}{\tau} (X(t) \mu) dt + \sigma dB(t)$
- Markovian property  $X(t_j^*)|X(t_{j-1}^*), \mu, \sigma, \tau \sim \mathcal{N}\left[\text{mean: } \mu + e^{-(t_j^* - t_{j-1}^*)/\tau} (X(t_{j-1}^*) - \mu), \text{ variance: } \frac{\tau\sigma^2}{2} (1 - e^{-2(t_j^* - t_{j-1}^*)/\tau})\right]$

$$\blacktriangleright \ p(\Delta, c) = p(\Delta)p(c) \propto I_{\{|\Delta| \in [0, (t_n-t_1)]\}}$$

#### BAYESIAN APPROACH: HYPER-PRIOR

- $\mu$  is a mean parameter of the underlying process
- $\blacktriangleright~\sigma$  is a scale parameter of the underlying process
- $\blacktriangleright \ \tau$  is a relaxation time of the underlying process
- Naively informative hyper-prior distribution:

$$p(\mu, \sigma^2, \tau) = p(\mu)p(\sigma^2)p(\tau) \propto rac{e^{-0.01/\sigma^2}}{(\sigma^2)^{1.01}} rac{e^{-1/\tau}}{\tau^2}$$

Flat on  $\mu$ , InvGamma(0.01, 0.01) on  $\sigma^2$ , and InvGamma(1, 1) on  $\tau$ 

### Full Posterior and Sampler based on ASIS

Suppose  $\theta_{hyp} \equiv (\mu, \sigma, \tau)$  and  $D_{obs} \equiv (\mathbf{x}(\mathbf{t}), \mathbf{y}(\mathbf{t}))$ Full Posterior:  $p(\mathbf{X}(\mathbf{t}^*), \Delta, c, \theta | D_{obs})$   $\propto L(\mathbf{X}(\mathbf{t}^*), \Delta, c) \cdot p(\mathbf{X}(\mathbf{t}^*), \Delta, c | \theta_{hyp}) \cdot p(\theta_{hyp})$ Likelihood Prior Hyper-prior

Metropolis-Hastings within Gibbs

- $p(\mathbf{X}(\mathbf{t}^*), \Delta | D_{obs}, c, \theta_{hyp})$
- $p(c|D_{obs}, \mathbf{X}(\mathbf{t}^*), \Delta, \theta_{hyp})$
- $p(\theta_{hyp}|D_{obs}, c, \mathbf{X}(\mathbf{t}^*), \Delta)$

Ancillarity-Sufficiency Interweaving Strategy (Yu and Meng, 2011)

- $p(\mathbf{X}(\mathbf{t}^*), \Delta | D_{obs}, c, \theta_{hyp})!$
- ▶ Interweaving  $p(c|D_{obs}, X(t^*), \Delta, \theta_{hyp})$  with  $p(c|D_{obs}, S(t^*), \Delta, \theta_{hyp})$
- $p(\theta_{hyp}|D_{obs}, c, \mathbf{X}(\mathbf{t}^*), \Delta)$

### EXAMPLE 1: SIMULATED DATA



Summary of posterior  $\Delta_{AB}$  with the blinded truth 30.98

Post. Mean	Post. Median	Post. SD	Half-length of 68% PI
30.94	30.89	0.694	0.423



## EXAMPLE 2: REAL DATA (Q0957+561)





Researchers	Method	$\hat{\Delta}_{AB}$	$SE(\hat{\Delta}_{AB})$
Oscoz et al. (1997)	Discrete cross-correlation & Dispersion	424	3
Serra-Ricart et al. (1999)	Cross-correlation functions	425	4
Tak et al. (?)	Bayesian	423.16	1.22

# EXAMPLE 2: REAL DATA (Q0957+561) (cont.)

Convergence Checks Each row: Traceplot, ACF, and histogram from the top Each column:  $\Delta, c, \mu, \log(\sigma), \log(\tau)$  from the left



#### DISCUSSION WITH REFERENCE

- Prior on Δ
- Quadruply-lensed quasar data
- Microlensing
- Reference
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  - B. Kelly, J. Bechtold, and A. Siemiginowska (2009) "Are the variations in quasar optical flux driven by thermal fluctuation?" The Astrophysical Journal, 698, 895 - 910.
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